

# 4769 Statistics 4

Q1				
(i)	$L = \frac{e^{-\theta}\theta^{x_1}}{x_1!} \cdots \frac{e^{-\theta}\theta^{x_n}}{x_n!} \left[ = \frac{e^{-n\theta}\theta^{\sum x_i}}{x_1!x_2!\cdots x_n!} \right]$ $\ln L = \text{const} - n\theta + \sum x_i \ln \theta$ $\frac{d \ln L}{d\theta} = -n + \frac{\sum x_i}{\theta} = 0$ $\Rightarrow \hat{\theta} = \frac{\sum x_i}{n} (= \bar{x})$ <p>Check this is a maximum</p> <p>e.g. <math>\frac{d^2 \ln L}{d\theta^2} = -\frac{\sum x_i}{\theta^2} &lt; 0</math></p>	M1 A1 M1 A1 A1 M1 A1	product form fully correct   CAO  	
				9
(ii)	$\lambda = P(X = 0) = e^{-\theta}$	B1		1
(iii)	We have $R \sim B(n, e^{-\theta})$ , so $E(R) = ne^{-\theta}$ $\text{Var}(R) = ne^{-\theta}(1 - e^{-\theta})$ $\tilde{\lambda} = \frac{R}{n}$ $\therefore E(\tilde{\lambda}) = e^{-\theta}$ i.e. unbiased $\text{Var}(\tilde{\lambda}) = \frac{e^{-\theta}(1 - e^{-\theta})}{n}$	M1 B1 B1 M1 A1 A1 A1	BEWARE PRINTED ANSWER	7

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(iv)	<p>Relative efficiency of <math>\tilde{\lambda}</math> wrt ML est</p> $= \frac{\text{Var(ML Est)}}{\text{Var}(\tilde{\lambda})}$ $= \frac{\theta e^{-2\theta}}{n} \cdot \frac{n}{e^{-\theta}(1-e^{-\theta})} = \frac{\theta}{e^\theta - 1}$ <p>Eg:- Expression is <math>\frac{\theta}{\theta + \frac{\theta^2}{2!} + \dots}</math></p> <p>always &lt; 1</p> <p>and this is <math>\approx 1</math> if <math>\theta</math> is small  <math>\approx 0</math> if <math>\theta</math> is large</p>	M1 M1 A1 M1 E1 E1 E1	any attempt to compare variances if correct BEWARE PRINTED ANSWER Allow statement that $\frac{\theta}{e^\theta - 1} \rightarrow 0$ as $\theta \rightarrow \infty$	7
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Q2			
(i)	$P(X = x) = q^{x-1} p$ $\text{Pgf } G(t) = E(t^X) = \sum_{x=1}^{\infty} pt^x q^{x-1}$ $= pt(1 + qt + q^2 t^2 + \dots)$ $= \underline{\underline{pt(1 - qt)^{-1}}}$  $\mu = G'(1) \quad \sigma^2 = G''(1) + \mu - \mu^2$ $G'(t) = pt(-1)(1 - qt)^{-2}(-q) + p(1 - qt)^{-1}$ $= pqt(1 - qt)^{-2} + p(1 - qt)^{-1}$  $\therefore G'(1) = pq(1 - q)^{-2} + p(1 - q)^{-1} = \frac{q}{p} + 1 = \frac{1}{\underline{\underline{p}}}$ $G''(t) = pqt(-2)(1 - qt)^{-3}(-q) + pq(1 - qt)^{-2} +$ $p(-1)(1 - qt)^{-2}(-q)$ $\therefore G''(1) = 2pq^2(1 - q)^{-3} + pq(1 - q)^{-2} + pq(1 - q)^{-2}$ $= \frac{2q^2}{p^2} + \frac{2q}{p}$ $\therefore \sigma^2 = \frac{2q^2}{p^2} + \frac{2q}{p} + \frac{1}{p} - \frac{1}{p^2} = \frac{2q^2 + 2pq + p - 1}{p^2}$ $= \frac{q}{p^2}(2q + 2p - 1) = \underline{\underline{\frac{q}{p^2}}}$	B1 M1 A1 A1 M1 A1 A1 A1 A1 M1 A1	FT into pgf only BEWARE PRINTED ANSWER [consideration of $ qt  < 1$ not required] for attempt to find $G'(t)$ and/or $G''(t)$ BEWARE PRINTED ANSWER For inserting their values BEWARE PRINTED ANSWER

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(ii)	$\left. \begin{array}{l} X_1 = \text{number of trials to first success} \\ X_2 = \text{ " " " " next " } \\ \vdots \\ \vdots \\ X_n = \text{ " " " " nth " } \end{array} \right\} \therefore Y = X_1 + X_2 + \dots + X_n$ <p style="text-align: right;">= total no of trials to the <math>n</math>th success</p> $\therefore \text{pgf of } Y = (\text{pgf of } X)^n = p^n t^n (1-qt)^{-n}$ $\mu_Y = n\mu_X = \frac{n}{p}$ $\sigma_Y^2 = n\sigma_X^2 = \frac{nq}{p^2}$	E1 E1	
(iii)	N(candidate's $\mu_Y$ , candidate's $\sigma_Y^2$ )	1	1
(iv)	<p>Y = no of tickets to be sold ~ random variable as in (ii) with <math>n = 140</math> and <math>p = 0.8</math></p> <p>~ Approx N( <math>\frac{140}{0.8} = 175</math>, <math>\frac{140 \times 0.2}{(0.8)^2} = 43.75</math> )</p> $P(Y \geq 160) \approx P(N(175, 43.75) > 159 \frac{1}{2})$ $= P(N(0,1) > -2.343)$ $= 0.9905$ <p>For any sensible discussion <u>in context</u> (eg groups of passengers <math>\Rightarrow</math> not indep.)</p>	E1 1 M1 A1 A1 E1 E1	<p>Do not award if cty corr absent or wrong, but FT if 160 used <math>\rightarrow</math> -2.268, 0.9884</p> <p>CAO</p>
Q3	$X = \text{amount of salt} \sim N(\mu[750], \sigma^2[20^2])$ Sample of $n=9$		
(i)	<p>Type I error: rejecting null hypothesis ... ... when it is true.</p> <p>Type II error: accepting null hypothesis ... ... when it is false.</p> <p>OC: P (accepting null hypothesis ... ... as a function of the parameter under investigation)</p>	B1 B1 B1 B1 B1 B1	<p>Allow B1 for <math>P(\text{rej } H_0 \text{ when true})</math></p> <p>Allow B1 for <math>P(\text{acc } H_0 \text{ when false})</math></p> <p>[ <math>P(\text{type II error} \mid \text{the true value of the parameter})</math> scores B1+B1 ]</p>
(ii)	<p>Reject if <math>\bar{x} &lt; 735</math> or <math>\bar{x} &gt; 765</math></p> $\alpha = P(\bar{X} < 735 \text{ or } \bar{X} > 765 \mid \bar{X} \sim N(750, \frac{20^2}{9}))$ $= P(Z < \frac{(735 - 750)3}{20} = -2.25$ $\text{or } Z > \frac{(765 - 750)3}{20} = 2.25)$ $= 2(1 - 0.9878) = 2 \times 0.0122 = 0.0244$ <p>This is the probability of rejecting good output and unnecessarily re-calibrating the machine – seems small [but not very small?]</p>	M1 A1 A1 A1 E1 E1	<p>Might be implicit</p> <p>CAO</p> <p>Accept any sensible comments</p>

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(iii)	<p>Accept if <math>735 &lt; \bar{x} &lt; 765</math>, and now <math>\mu = 725</math>.</p> $\beta = P(735 < \bar{X} < 765 \mid \bar{X} \sim N(725, \frac{20^2}{9}))$ $= P(1.5 < Z < 6)$ $= 1 - 0.9332 = \underline{\underline{0.0668}}$ <p>This is the probability of accepting output and carrying on when in fact <math>\mu</math> has slipped to 725 – small[-ish?]</p>	M1 A1 A1 A1 E1 E1	<p>might be implicit</p> <p>CAO If upper limit 765 not considered, maximum 2 of these 4 marks. If <math>\Phi(6)</math> not considered, maximum 3 out of 4. accept sensible comments</p>	6
(iv)	$OC = P(735 < \bar{X} < 765 \mid \bar{X} \sim N(\mu, \frac{20^2}{9}))$ $= \Phi\left(\frac{(765 - \mu)3}{20}\right) - \Phi\left(\frac{(735 - \mu)3}{20}\right)$ <p>" <math>\Phi - \Phi</math>"</p> <p><math>\mu = 720: \Phi(6.75) - \Phi(2.25) = 1 - 0.9878 = 0.0122</math>  <math>730: 5.25 - 0.75 = 1 - 0.7734 = 0.2266</math>  <math>740: 3.75 - 0.75 = 1 - (1 - 0.7734) = 0.7734</math></p> <p>750: similarly or by write-down from part (ii)  [FT]: 0.9756</p> <p>760, 770, 780 by symmetry  [FT]: 0.7734, 0.2266, 0.0122</p>	M1 M1 A1 1 1 1	<p>both correct</p> <p>if any two correct</p>	
Q4				6
(i)	$x_{ij} = \mu + \alpha_i + e_{ij}$ $\mu$ = population ... .. grand mean for whole experiment $\alpha_i$ = population ... .. mean by which $i$ th treatment differs from $\mu$ $e_{ij}$ are experimental errors... $\sim \text{ind } N(0, \sigma^2)$	1 1 1 1 1 1 3	Allow "uncorrelated" 1 for ind N; 1 for 0; 1 for $\sigma^2$ .	9
(ii)	<p>Totals are 240, 246, 254, 264, 196 each from sample of size 5  Grand total 936</p> <p>"Correction factor" CF = <math>\frac{936^2}{20} = 43804.8</math></p> <p>Total SS = 44544 - CF = 739.2</p>			

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	Between contractors SS = $\frac{240^2}{5} + \dots + \frac{196^2}{5} - CF = 44209.6 - CF = 404.8$ Residual SS ( by subtraction) = $739.2 - 404.8 = 334.4$	M1 M1 A1	For correct methods for any two, if each calculated SS is correct.																					
	<table border="1" style="margin-left: auto; margin-right: auto;"> <thead> <tr> <th>Source of Variation</th> <th>SS</th> <th>df</th> <th>MS</th> <th>MS ratio</th> </tr> </thead> <tbody> <tr> <td>Between Contractors</td> <td>404.8</td> <td>3</td> <td>134.93</td> <td>6.456</td> </tr> <tr> <td>Residual</td> <td>334.4</td> <td>16</td> <td>20.9</td> <td></td> </tr> <tr> <td>Total</td> <td>739.2</td> <td>19</td> <td></td> <td></td> </tr> </tbody> </table>	Source of Variation	SS	df	MS	MS ratio	Between Contractors	404.8	3	134.93	6.456	Residual	334.4	16	20.9		Total	739.2	19			M1 M1 1 A1	CAO	
Source of Variation	SS	df	MS	MS ratio																				
Between Contractors	404.8	3	134.93	6.456																				
Residual	334.4	16	20.9																					
Total	739.2	19																						
	Refer to $F_{3,16}$	1	NO FT IF WRONG																					
	Upper 5% point is 3.24	1	NO FT IF WRONG																					
	Significant	1																						
	Seems performances of contractors are not all the same	1																						
				12																				
(iii)	Randomised blocks  Description	B1 E1 E1	Take the subject areas as "blocks", ensure each contractor is used at least once in each block	3																				